Data Science and Advanced Programming — Lecture 11

Unsupervised Machine Learning

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Today's Roadmap

- 1. k-Means
- 2. Evaluating Clusterings (read at home here for completeness)
- 3. Gaussian Mixture Models
- 4. Principal Component Analysis (PCA)
- 5. Expectation Maximization (read at home here for completeness)
- 6. Hierarchical Clustering
- 7. Density-based Clustering (cont'd)

Learning Parameters of Probability Distributions

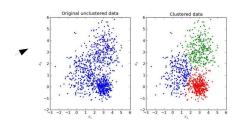
- ▶ In many settings not all variables are observed (labeled) in the training data $x_i = (x_i, h_i)$
- ▶ e.g. Speech recognition: have speech signals, but not phoneme labels.
- e.g. object recognition: have object labels (car, bike), but not part labels (wheel, door, seat).
- ► Unobserved variables are called **LATENT VARIABLES**.

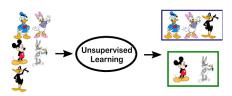


Recall — Unsupervised Learning

Learning "what normally happens".

- ► No output.
- ► Clustering: Grouping similar instances.
- ► Example applications:
 - ► Customer segmentation.
 - Image compression: Color quantization.
 - ► Bioinformatics: Learning motifs.





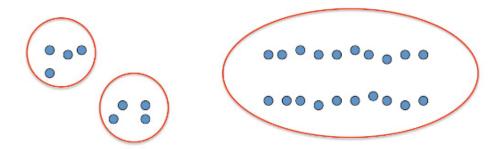
Motivation — Clustering

- ► Clustering, as a kind of **unsupervised learning**, aims at grouping data points into clusters.
- ► Intuition: Data points within
 - ► the same cluster should be close to each other
 - ▶ different clusters should be far apart from each other
- ► Applications:
 - ► segmentation of customers (e.g., for marketing campaigns)
 - ▶ organization/exploration of data (e.g., search results)
 - detection of outliers data points

Clustering: Basic idea

► Basic idea: group together similar instances

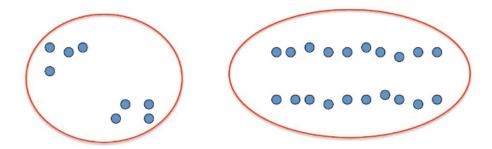
► Example: 2D point patterns



Clustering: Basic idea

► Basic idea: group together similar instances

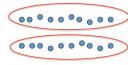
► Example: 2D point patterns



Clustering: Basic idea

- Basic idea: group together similar instances
- ► Example: 2D point patterns



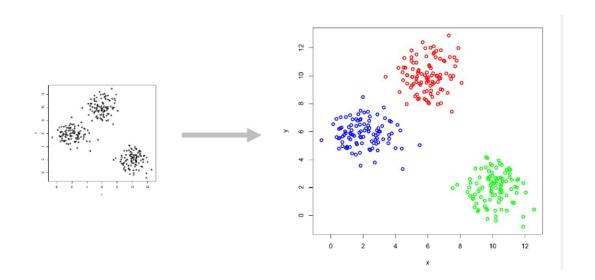


- What could similar mean?
- One option: small Euclidean distance (squared)

$$\operatorname{dist}(\vec{x}, \vec{y}) = \|\vec{x} - \vec{y}\|_2^2$$

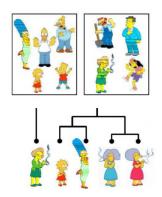
► Clustering results are crucially dependent on the measure of similarity (or distance) between "points" to be clustered

Clustering: Basic idea in color



Clustering Algorithms

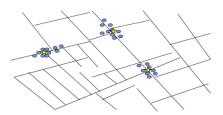
- ► Partition algorithms (Flat)
 - ► K-Means
 - ► Mixture of Gaussians
 - ▶ ...
- ► Hierarchical algorithms
 - ► Bottom-up agglomerative
 - ► Top down divisive



First (?) Application of Clustering

- ▶ John Snow, a London physician plotted the location of **cholera deaths** on a map during an outbreak in the 1850s.
- ► The locations indicated that cases were clustered around certain intersections where there were polluted wells thus exposing both the problem and the solution.

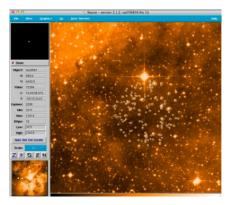




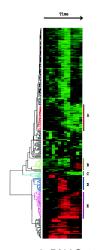
From: Nina Mishra HP Labs

Clustering Example: Astronomy

SkyCat (http://www.eso.org/sci/observing/tools/skycat.html): Clustered 2×10^9 sky objects into stars, galaxies, quasars, etc. based on radiation emitted in different spectrum bands



Another Clustering Example: Genetics

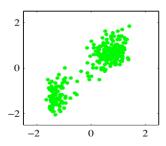


Eisen et al, PNAS 1998

1. k-Means – Unsupervised ML

Bishop, Chapter 9

- ► We will start with an unsupervised learning (clustering) problem:
- ▶ Given a dataset $\{x_1, \dots, x_N\}$ each $x_i \in \mathbb{R}^D$ partition the dataset into K clusters (e.g. healthy / sick patients).
- ► Intuitively, a cluster is a group of points, which are close together and far from others.



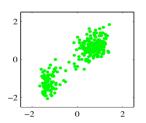
Distortion Measure

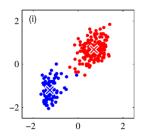
- ► Formally, introduce prototypes (or cluster centers) $\mu_k \in \mathbb{R}^D$
- ▶ Use binary r_{nk} , 1 if point n is in cluster k, 0 otherwise (1-of- K coding scheme again)
- Find $\{\mu_k\}$, $\{r_{nk}\}$ to minimize distortion measure:

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \| \mathbf{x}_{n} - \boldsymbol{\mu}_{k} \|^{2}$$

e.g. two clusters k = 1, 2:

$$J = \sum_{\mathbf{x}_n \in C_1} \|\mathbf{x}_n - \mu_1\|^2 + \sum_{\mathbf{x}_n \in C_2} \|\mathbf{x}_n - \mu_2\|^2$$





Minimizing Distortion Measure

► Minimizing *J* directly is hard

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \| \mathbf{x}_{n} - \boldsymbol{\mu}_{k} \|^{2}$$

- ► However, two things are easy
 - ▶ If we know μ_k , minimizing J wrt r_{nk}
 - ▶ If we know r_{nk} , minimizing J wrt μ_k
- ► This suggests an iterative procedure
 - ▶ Start with initial guess for μ_k
 - ► Iteration of two steps:
 - ightharpoonup Minimize J wrt r_{nk}
 - ► Minimize J wrt μ_k
 - Rinse and repeat until convergence

Determining Membership Variables

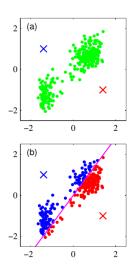
➤ Step 1 in an iteration of K-means is to minimize distortion measure J wrt. cluster membership variables *r*_{nk}

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_{n} - \boldsymbol{\mu}_{k}\|^{2}$$

▶ Terms for different data points x_n are independent, for each data point set r_{nk} to minimize

$$\sum_{k=1}^K r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

► Simply set $r_{nk} = 1$ for the cluster center μ_k with smallest distance.



Determining Cluster Centers

▶ Step 2: fix r_{nk} , minimize J wrt the cluster centers μ_k

$$J = \sum_{k=1}^{K} \sum_{n=1}^{N} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 \text{ switch order of sums}$$

- ▶ So we can minimize wrt each μ_k separately
- Take derivative, set to zero:

$$2\sum_{n=1}^{N} r_{nk} (\mathbf{x}_n - \boldsymbol{\mu}_k) = 0$$

$$\Leftrightarrow \boldsymbol{\mu}_k = \frac{\sum_n r_{nk} \mathbf{x}_n}{\sum_n r_{nk}}$$

i.e. mean of datapoints x_n assigned to cluster $k \ (\rightarrow \text{"k-Means"})$

k-Means Algorithm

- ▶ Start with an initial guess for μ_k
- ► Iteration of two steps:
 - 1. Minimize J wrt r_{nk}
 - ► Assign points to nearest cluster center
 - 2. Minimize J wrt μ_k
 - ► Set cluster center as average of points in cluster
- ► Rinse and repeat until convergence

Old Faithful Dataset

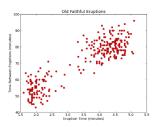
https://www.stat.cmu.edu/~larry/all-of-statistics/=data/faithful.dat

```
Description: (From R manual):

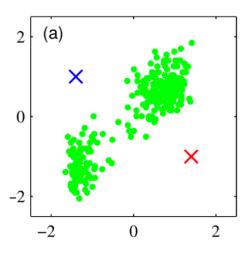
Waiting time between eruptions and the duration of the eruption for the Old Faithful geyser in Yellowstone National Park, Wyoming, USA.

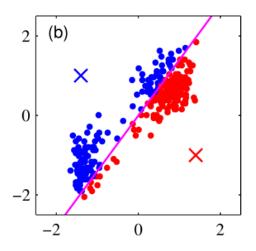
A data frame with 272 observations on 2 variables.

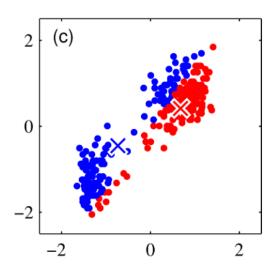
eruptions numeric Eruption time in mins waiting numeric Waiting time to next eruption
```

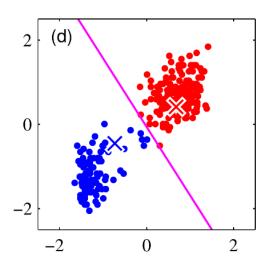


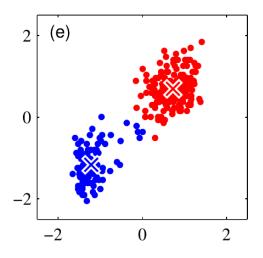


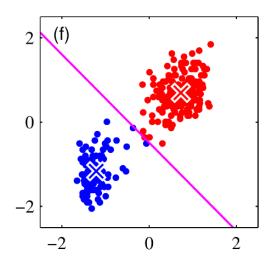


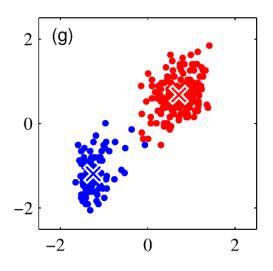


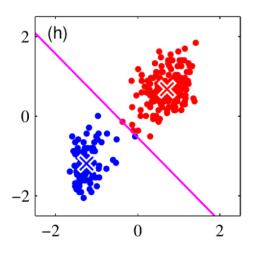


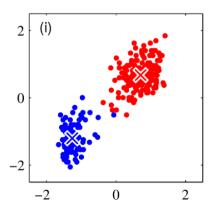






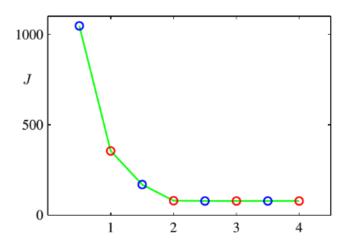






Next step doesn't change membership — stop

Cost function J



k-means Convergence

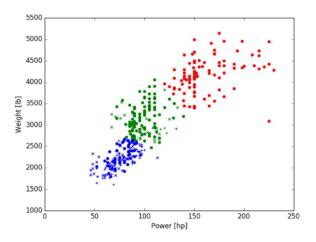
- ► Repeat steps until no change in cluster assignments.
- ► For each step, value of J either goes down, or we stop.
- ► Finite number of possible assignments of data points to clusters, so we are guaranteed to converge eventually.
- Note it may be a local maximum rather than a global maximum to which we converge.

Clustering Cars based on Power and Weight

demo/k_means_car.py

```
import numpy as np
import pandas as pd
from sklearn.model selection import train_test_split
from sklearn.preprocessing import MinMaxScaler
from sklearn.cluster import KMeans
import matplotlib.pyplot as plt
# load data
cars = pd.read_csv('auto-mpg.data.txt',header=None, sep='\s+')
# extract power and weight as data matrix X
X = cars.iloc[:, [3,4]].values
# extract origin as target value y
v = cars.iloc[:, 7].values
# normalize data
min_max_scaler = MinMaxScaler()
min_max_scaler.fit(X) # determine min and max
X_normalized = min_max_scaler.transform(X)
# applu k-Means
km = KMeans(n clusters=3, random state=0).fit(X normalized)
# plot cars
# U.S. : o / Europe: x / Japan : +
m = ['o' \text{ if } o==1 \text{ else } 'x' \text{ if } o==2 \text{ else } '+' \text{ for } o \text{ in } y]
# Cluster 1 : red / Cluster 2 : blue / Cluster 3 : green
c = ['red' if l==0 else 'blue' if l==1 else 'green' for l in km.labels_]
for i in range(0.len(X)):
    plt.scatter(X[i,0], X[i,1], color=c[i], marker=m[i])
plt.xlabel('Power [hp]')
plt.ylabel('Weight [lb]')
plt.show()
```

Clustering Cars based on Power and Weight: Plot



2. Evaluating Clustering: Some Notation

Read at home — here for completeness

- ► Consider a set of data points $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ $\mathbf{x}_i \in \mathbb{R}^m$
- ► Objective: Determine clustering (also: grouping, partitioning)

$$C = \{C_1, \dots, C_k\}$$
 with $C_i \subseteq D$

such that

- ▶ clusters are disjoint $\forall i \neq j : C_i \cap C_i = \emptyset$
- ► each data point is assigned to a cluster

$$\bigcup_{C_i\in\mathcal{C}}C_i=\mathcal{D}$$

Evaluating Clustering

- ► How can we evaluate the quality of a clustering computed?
- ► External measures assume that ideal clustering is known (e.g., class labels assigned to data points)

$$\mathcal{I} = \left\{ \emph{I}_1, \ldots, \emph{I}_{|\mathcal{I}|}
ight\} \quad \text{ with } \quad \emph{I}_i \subseteq \mathcal{D}$$

► Internal measures assume no knowledge of ideal clustering (i.e., we only know the data points and the clustering)

Purity

Purity of a cluster is the fraction of data points therein that belongs to the dominant cluster from the ideal clustering

$$\operatorname{purity}\left(C_{i}\right) = \frac{1}{|C_{i}|} \max_{I_{j} \in \mathcal{I}} |C_{i} \cap I_{j}|$$

► Purity of a clustering is then the weighted average of the purity values of its clusters

$$\operatorname{purity}(\mathcal{C}) = \sum_{C_i \in \mathcal{C}} \frac{|C_i|}{n} \operatorname{purity}(C_i)$$

Purity







$$\operatorname{purity}(\mathit{C}_1) = \frac{2}{3}$$

$$purity(C_2) = \frac{2}{4}$$

$$\operatorname{purity}(C_3) = \frac{1}{2}$$

purity(
$$\mathcal{C}$$
) = $\frac{3}{9} \cdot \frac{2}{3} + \frac{4}{9} \cdot \frac{2}{4} + \frac{2}{9} \cdot \frac{1}{2} \approx 0.56$

BetaCV

▶ BetaCV, as an internal measure, considers the ratio of average distances between pairs of points within the same or different clusters

$$\mathsf{BetaCV}(\mathcal{C}) = rac{W_\mathsf{in} \ / N_\mathsf{in}}{W_\mathsf{out} \ / N_\mathsf{out}}$$

ightharpoonup with $N_{\rm in}$ and $N_{\rm out}$ as pairs of data points within the same or within different clusters

$$N_{\mathsf{in}} \ = rac{1}{2} \sum_{C_i \in \mathcal{C}} |C_i| \, (|C_i| - 1) \quad N_{\mathsf{out}} \ = rac{1}{2} \sum_{C_i, C_j \in \mathcal{C}, C_i
eq C_j} |C_i| \, |C_j|$$

BetaCV

▶ BetaCV, as an internal measure, considers the ratio of average distances between pairs of points within the same or different clusters

$$\mathsf{BetaCV}(\mathcal{C}) = rac{W_\mathsf{in} \ / N_\mathsf{in}}{W_\mathsf{out} \ / N_\mathsf{out}}$$

ightharpoonup and W_{out} as the total distance of pairs of data points within the same or within different clusters

$$W_{\text{in}} = \frac{1}{2} \sum_{C_i \in \mathcal{C}} \sum_{\mathbf{x}, \mathbf{y} \in C_i} d(\mathbf{x}, \mathbf{y}) \quad W_{\text{out}} = \frac{1}{2} \sum_{C_i, C_j \in \mathcal{C}} \sum_{\mathbf{x} \in C_i} \sum_{\mathbf{y} \in C_j} d(\mathbf{x}, \mathbf{y})$$

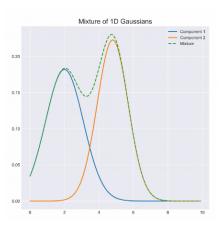
Dunn Index

▶ Dunn Index, as another **internal measure**, compares the minimal distance between any pair of data points from different clusters against the maximal distance between any pair of data points from the same cluster.

$$\mathsf{DunnIndex}(\mathcal{C}) = \frac{\min\limits_{\mathbf{x} \in C_i, \mathbf{y} \in C_j, C_i \neq C_j} d(\mathbf{x}, \mathbf{y})}{\max\limits_{\mathbf{x} \in C_i, \mathbf{y} \in C_i} d(\mathbf{x}, \mathbf{y})}$$

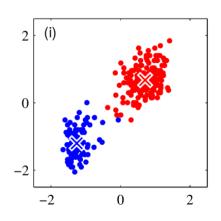
3. Gaussian Mixture Models

See Bishop (2006), Chapter 9; Murphy (2012), Chapter 11

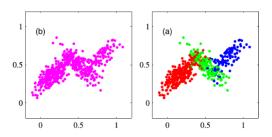


Hard Assignment versus Soft Assignment

- ► In the K-means algorithm, a hard assignment of points to clusters is made.
- ► However, for points near the decision boundary, this may not be such a good idea.
- ► Instead, we could think about making a **soft** assignment of points to clusters.



Gaussian Mixture Models

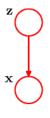


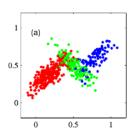
- ► The Gaussian mixture model (or Mixture of Gaussians MoG) models the data as a combination of Gaussians.
- ► Above shows a dataset generated by drawing samples from three different Gaussians.

$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k
ight). \quad p\left(z_k = 1
ight) = \pi_k$$

A Generative Model

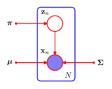
- ► The mixture of Gaussians is a generative model.
- ▶ To generate a data point x_n , we first generate a value for a discrete variable $z_n \in \{1, ..., K\}$
- lacktriangle We then generate a value $oldsymbol{x}_n \sim \mathcal{N}\left(oldsymbol{x} \mid oldsymbol{\mu}_k, \Sigma_k
 ight)$ for the corresponding Gaussian

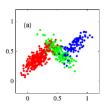




A Graphical Model

- ightharpoonup Note z_n is a latent variable, unobserved.
- ▶ Need to give conditional distributions $p(z_n)$ and $p(x_n \mid z_n)$
- ▶ The one-of-K representation is helpful here: $z_{nK} \in \{0,1\}$, $z_n = (z_{n1}, \dots, z_{nK})$

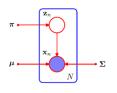


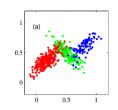


Graphical Model — Latent Component Variable

- ▶ Use a Bernoulli distribution for $p(z_n)$
 - ▶ i.e. $p(z_{nk} = 1) = \pi_k$
 - ▶ Parameters to this distribution $\{\pi_K\}$
- ▶ Must have $0 \le \pi_k \le 1$ and $\sum_{k=1}^K \pi_k = 1$
- $\triangleright p(\mathbf{z}_n) = \prod_{k=1}^K \pi_k^{\mathbf{z}_{nk}}$

Graphical Model – Observed Variable

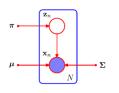


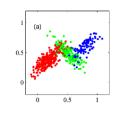


- ▶ Use a Gaussian distribution for $p(x_n | z_n)$
- ▶ Parameters to this distribution $\{\mu_k, \Sigma_k\}$

$$egin{aligned}
ho\left(\mathbf{x}_n\mid \mathbf{z}_{nk}=1
ight) &= \mathcal{N}\left(\mathbf{x}_n\mid oldsymbol{\mu}_k, oldsymbol{\Sigma}_k
ight) \
ho\left(\mathbf{x}_n\mid oldsymbol{z}_n
ight) &= \prod_{k=1}^K \mathcal{N}\left(\mathbf{x}_n\mid oldsymbol{\mu}_k, oldsymbol{\Sigma}_k
ight)^{oldsymbol{z}_{nk}} \end{aligned}$$

A Graphical Model – Joint Distribution





► The full **joint distribution** is given by:

$$p(\mathbf{x}, \mathbf{z}) = \prod_{n=1}^{N} p(\mathbf{z}_n) p(\mathbf{x}_n \mid \mathbf{z}_n)$$

$$= \prod_{n=1}^{N} \prod_{k=1}^{K} \pi_k^{\mathbf{z}_{nk}} \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{\mathbf{z}_{nk}}$$

Marginal over Observed (MoG) Variables

▶ The marginal distribution $p(x_n)$ for this model is:

$$\begin{aligned} \rho(\mathbf{x}_n) &= \sum_{\mathbf{z}_n} \rho(\mathbf{x}_n, \mathbf{z}_n) = \sum_{\mathbf{z}_n} \rho(\mathbf{z}_n) \, \rho(\mathbf{x}_n \mid \mathbf{z}_n) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{aligned}$$

► A mixture of Gaussians

MoG Conditional over Latent Variable

- ▶ The conditional $p(z_{nk} = 1 \mid x_n)$ will play an important role for learning
- ▶ It is denoted by $\gamma(z_{nk})$ can be computed as:

$$\gamma(z_{nk}) \equiv p(z_{nk} = 1 \mid \mathbf{x}_n) = \frac{p(z_{nk} = 1) p(\mathbf{x}_n \mid z_{nk} = 1)}{\sum_{j=1}^{K} p(z_{nj} = 1) p(\mathbf{x}_n \mid z_{nj} = 1)}$$
$$= \frac{\pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

 $ightharpoonup \gamma(z_{nk})$ is the responsibility of component k for datapoint n

MoG Learning

- ▶ Given a set of observations $\{x_1, \ldots, x_N\}$, without the latent
- \triangleright variables z_n , how can we learn the parameters?
- ▶ Model parameters are $\theta = \{\pi_k, \mu_k, \Sigma_k\}$
- ► Answer will be similar to k-means:
 - ightharpoonup If we know the latent variables z_n , fitting the Gaussians is easy
 - ▶ If we know the Gaussians μ_k, Σ_k , finding the latent variables is easy
- ightharpoonup Rather than latent variables, we will use responsibilities $\gamma(z_{nk})$

MoG Maximum Likelihood Learning

- ▶ Given a set of observations $\{x_1, \ldots, x_N\}$, without the latent variables z_n , how can we learn the parameters?
- ▶ Model parameters are $\theta = \{\pi_k, \mu_k, \Sigma_k\}$
- ► We can use the maximum likelihood criterion:

$$egin{aligned} eta_{\mathit{ML}} &= \arg\max_{m{ heta}} \prod_{n=1}^{N} \sum_{k=1}^{K} \pi_{k} \mathcal{N}\left(m{x}_{n} \mid m{\mu}_{k}, m{\Sigma}_{k}
ight) \ &= \arg\max_{m{ heta}} \sum_{n=1}^{N} \log \left\{ \sum_{k=1}^{K} \pi_{k} \mathcal{N}\left(m{x}_{n} \mid m{\mu}_{k}, m{\Sigma}_{k}
ight)
ight\} \end{aligned}$$

► Unfortunately, closed-form solution not possible this time - log of sum rather than log of product

MoG Maximum Likelihood Learning - Problem

► Maximum likelihood criterion, 1-D:

$$\theta_{ML} = \arg\max_{\theta} \sum_{n=1}^{N} \log \left\{ \sum_{k=1}^{K} \pi_{k} \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{ -\left(x_{n} - \mu_{k}\right)^{2} / \left(2\sigma^{2}\right) \right\} \right.$$

• Suppose we set $\mu_k = x_n$ for some k and n, then we have one term in the sum:

$$\pi_{k} \frac{1}{\sqrt{2\pi}\sigma_{k}} \exp\left\{-\left(x_{n} - \mu_{k}\right)^{2} / \left(2\sigma^{2}\right)\right\}$$
$$= \pi_{k} \frac{1}{\sqrt{2\pi}\sigma_{k}} \exp\left\{-\left(0\right)^{2} / \left(2\sigma^{2}\right)\right\}$$

- ▶ In the limit as $\sigma_k \to 0$, this goes to ∞
- ▶ So ML solution is to set some $\mu_k = x_n$, and $\sigma_k = 0$!

ML for Mixture of Gaussians

- ► Keeping this problem in mind, we will develop an algorithm for ML estimation of the parameters for a MoG model
- ► Search for a local optimum.
- ► Consider the log-likelihood function

$$\ell(oldsymbol{ heta}) = \sum_{n=1}^{N} \log \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}\left(oldsymbol{x}_n \mid oldsymbol{\mu}_k, oldsymbol{\Sigma}_k
ight)
ight\}$$

▶ We can try taking derivatives and setting to zero, even though no closed form solution exists.

Maximizing Log-Likelihood - Means

$$\ell(\boldsymbol{\theta}) = \sum_{n=1}^{N} \log \left\{ \sum_{k=1}^{K} \pi_{k} \mathcal{N} \left(\boldsymbol{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k} \right) \right\}$$

$$\frac{\partial}{\partial \mu_{k}} \ell(\boldsymbol{\theta}) = \sum_{n=1}^{N} \frac{\pi_{k} \mathcal{N} \left(\boldsymbol{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k} \right)}{\sum_{j} \pi_{j} \mathcal{N} \left(\boldsymbol{x}_{n} \mid \boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j} \right)} \boldsymbol{\Sigma}_{k}^{-1} \left(\boldsymbol{x}_{n} - \boldsymbol{\mu}_{k} \right)$$

$$= \sum_{n=1}^{N} \gamma \left(\boldsymbol{z}_{nk} \right) \boldsymbol{\Sigma}_{k}^{-1} \left(\boldsymbol{x}_{n} - \boldsymbol{\mu}_{k} \right)$$

- Setting derivative to 0 , and multiply by Σ_k

$$\sum_{n=1}^{N} \gamma(z_{nk}) \mu_{k} = \sum_{n=1}^{N} \gamma(z_{nk}) x_{n}$$

$$\Leftrightarrow \mu_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) x_{n} \text{ where } N_{k} = \sum_{n=1}^{N} \gamma(z_{nk})$$

Maximizing Log-Likelihood: Means and Covariances

Note that the mean μ_k is a weighted combination of points x_n , using the responsibilities $y(z_{nk})$ for the cluster k

$$\mu_k = \frac{1}{N_k} \sum_{n=1}^{N} \gamma(z_{nk}) x_n$$

- $N_k = \sum_{n=1}^N \gamma(z_{nk})$ is the effective number of points in the cluster
- \blacktriangleright A similar result comes from taking derivatives wrt. the covariance matrices Σ_k :

$$\mathbf{\Sigma}_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma\left(z_{nk}\right) \left(\mathbf{x}_{n} - \mathbf{\mu}_{k}\right) \left(\mathbf{x}_{n} - \mathbf{\mu}_{k}\right)^{T}$$

Maximizing Log-Likelihood: Mixing Coefficients

- \blacktriangleright We can also maximize wrt. the mixing coefficients π_k
- ▶ Note there is a constraint that $\sum_k \pi_k = 1$
- ► Use Lagrange multipliers
- ▶ End up with: $\pi_k = \frac{N_k}{N}$ average responsibility that component k takes.

Three Parameter Types and Three Equations

► These three equations a solution does not make

$$egin{aligned} oldsymbol{\mu}_k &= rac{1}{N_k} \sum_{n=1}^N \gamma\left(z_{nk}
ight) oldsymbol{x}_n \ oldsymbol{\Sigma}_k &= rac{1}{N_k} \sum_{n=1}^N \gamma\left(z_{nk}
ight) \left(oldsymbol{x}_n - oldsymbol{\mu}_k
ight) \left(oldsymbol{x}_n - oldsymbol{\mu}_k
ight)^T \ \pi_k &= rac{N_k}{N} \end{aligned}$$

- ▶ All depend on $\gamma(z_{nk})$, which depends on all 3!
- ► But an iterative scheme can be used

EM for Mixtures of Gaussians

- ► Initialize parameters, then iterate:
 - ► **E step**: Calculate responsibilities using current parameters

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

▶ **M step**: Re-estimate parameters using these $y(z_{nk})$

$$egin{aligned} oldsymbol{\mu}_k &= rac{1}{N_k} \sum_{n=1}^N \gamma\left(z_{nk}
ight) oldsymbol{x}_n \ oldsymbol{\Sigma}_k &= rac{1}{N_k} \sum_{n=1}^N \gamma\left(z_{nk}
ight) \left(oldsymbol{x}_n - oldsymbol{\mu}_k
ight) \left(oldsymbol{x}_n - oldsymbol{\mu}_k
ight)^T \ \pi_k &= rac{N_k}{N} \end{aligned}$$

- ► This algorithm is known as the expectation-maximization algorithm (EM)
 - Next we describe its general form, why it works, and why it's called EM (but first an example)

The Likelihood

- ► The form of the Gaussian mixture distribution is governed by the parameters π , μ and Σ , where we have used the notation $\pi \equiv \{\pi_1, \dots, \pi_k\}$, $\mu \equiv \{\mu_1, \dots, \mu_k\}$, $\Sigma \equiv \{\Sigma_1, \dots \Sigma_k\}$.
- ▶ One way to set the values of these parameters is to use maximum likelihood.
- ► The log of the likelihood function is given by

$$\ln p(\mathbf{X} \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_{k} \mathcal{N} \left(\mathbf{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k} \right) \right\}$$

where
$$X = \{x_1, ..., x_N\}$$
.

Problems with optimizing the likelihood

- ► The situation is now much more complex than with a single Gaussian, due to the presence of the summation over k inside the logarithm.
- ► As a result, the maximum likelihood solution for the parameters no longer has a closed-form analytical solution.
- ► One approach to maximizing the likelihood function is to use iterative numerical optimization techniques.
- ► Gradient methods could be used but are painful to implement. ⇒ Non-convex optimization problem! (multiple optima possible)

Example in one dimension

- ightharpoonup Observations $x_1 \dots x_n$
- ► K = 2 Gaussians with unknown μ, σ^2
- ► Estimation trivial if we know the source of each observation

$$\mu_b = \frac{x_1 + x_2 + \dots + x_{n_b}}{n_b}$$

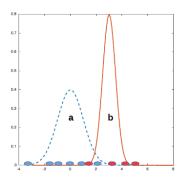
$$\sigma_b^2 = \frac{(x_1 - \mu_1)^2 + \dots + (x_n - \mu_n)^2}{n_b}$$

Example in one dimension

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Example: Expectation Maximization in 1d (II)

- ► What if we don't know the source?
- ► If we knew parameters of the Gaussians (μ, σ^2)



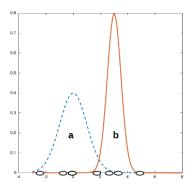
Example: Expectation Maximization in 1d (II)

- ▶ What if we don't know the source?
- ▶ If we knew parameters of the Gaussians (μ, σ^2)

 \rightarrow can guess whether point is more likely to be *a* or *b*.

$$P(b \mid x_i) = \frac{P(x_i \mid b) P(b)}{P(x_i \mid b) P(b) + P(x_i \mid a) P(a)}$$

$$P(x_i \mid b) = \frac{1}{\sqrt{2\pi\sigma_b^2}} \exp\left(-\frac{(x_i - \mu_b)^2}{2\sigma_b^2}\right)$$



EM Algorithm (in 1d)

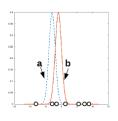
A fundamental problem:

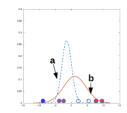
- we need (μ_a, σ_a^2) and (μ_b, σ_b^2) to guess the source of the points.
- we need to know the source to estimate (μ_a, σ_a^2) and (μ_b, σ_b^2) .

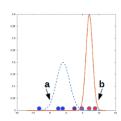
EM algorithm:

- 1. **Start** with two randomly placed Gaussians (μ_a, σ_a^2) and (μ_b, σ_b^2) .
- 2. **E(xpectation) step:**
 - ▶ for each point: $P(b \mid x_i) = \text{does it look like it came from } b$?
- M(aximization)-step:
 - ▶ adjust (μ_a, σ_a^2) and (μ_b, σ_b^2) to fit points assigned to them.
- 4. Iterate until convergence.

EM in 1d







$$P(x_{i} \mid b) = \frac{1}{\sqrt{2\pi\sigma_{b}^{2}}} \exp\left(-\frac{(x_{i} - \mu_{b})^{2}}{2\sigma_{b}^{2}}\right) \qquad \sigma_{b}^{2} = \frac{b_{1}x_{1} + b_{2}x_{2} + \dots + b_{n}x_{n_{b}}}{b_{1} + b_{2} + \dots + b_{n}} \rightarrow \text{We could also es}$$

$$\sigma_{b}^{2} = \frac{b_{1}(x_{1} - \mu_{1})^{2} + \dots + b_{n}(x_{n} - \mu_{n})^{2}}{b_{1} + b_{2} + \dots + b_{n}} \rightarrow \text{We could also es}$$

$$b_{i} = P(b \mid x_{i}) = \frac{P(x_{i} \mid b) P(b)}{P(x_{i} \mid b) P(b) + P(x_{i} \mid a) P(a)} \xrightarrow{a_{1}x_{1} + a_{2}x_{2} + \dots + a_{n}x_{n_{n}}} \qquad P(b) = (b_{1} + b_{2} + \dots + b_{n})$$

$$a_{i} = P(a \mid x_{i}) = 1 - b_{i} \qquad \sigma_{a}^{2} = \frac{a_{1}(x_{1} - \mu_{1})^{2} + \dots + a_{n}(x_{n} - \mu_{n})^{2}}{a_{1} + a_{2} + \dots + a_{n}}$$

 \rightarrow We could also estimate priors:

$$P(b) = (b_1 + b_2 + \dots b_n) / n$$

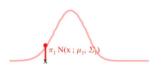
 $P(a) = 1 - P(b)$

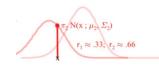
EM in the multidimensional case

- ► Start with parameters describing each cluster
- ► Mean μ_c , Covariance Σ_c , "size" π_c
- ► E-step ("Expectation"):
 - ightharpoonup For each observation/point x_i
 - Compute "r_{ic}", the probability that it belongs to cluster c.
 - ► Compute its probability under model c.
 - ► Normalize to sum to one (over clusters c).

$$r_{ic} = \frac{\pi_c \mathcal{N}\left(x_i; \mu_c, \Sigma_c\right)}{\sum_{c'} \pi_{c'} \mathcal{N}\left(x_i; \mu_{c'}, \Sigma_{c'}\right)}$$

- ► If *x_i* is very likely under the c-th Gaussian, it gets high weight.
- ▶ Denominator just makes r's sum to one.





EM in the multidimensional case

- ► M-step ("Maximization step"):
 - For each cluster (Gaussian) z = c
 - ► Update its parameters using the (weighted) data points

$$N_c = \sum_i r_{ic}$$
 Total responsibility allocated to cluster c $\pi_c = \frac{N_c}{N}$ Fraction of total assigned to cluster c $\mu_c = \frac{1}{N_c} \sum_i r_{ic} x_i$ $\Sigma_c = \frac{1}{N_c} \sum_i r_{ic} (x_i - \mu_c)^T (x_i - \mu_c)$

Weighted mean of assigned data

weighted covariance of assigned data (use new weighted means here)

Expectation-Maximization: Summary

► Likelihood of the data

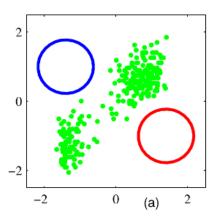
$$P(x_1,...,x_N) = \prod_{i=1}^{N} \sum_{k=1}^{K} P(x_i \mid k) P(k)$$

► Each step increases the log-likelihood of our model

$$\ln p(\mathbf{X} \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_{k} \mathcal{N} \left(\mathbf{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k} \right) \right\}$$

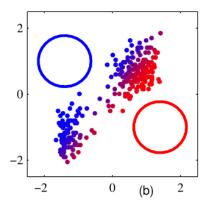
- ► Iterate until convergence
 - ► Convergence guaranteed another ascent method.
- Cannot discover k.

MoG EM — Example

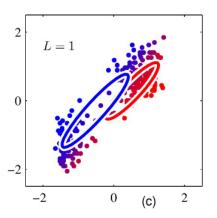


- ► Same initialization as with K-means before
- ► Often, K-means is actually used to initialize EM

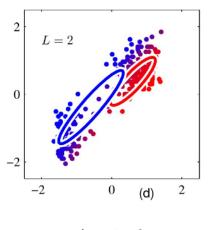
MoG EM — Example



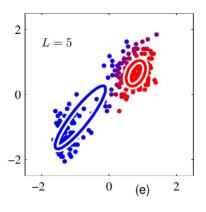
Calculate responsibilities $\gamma\left(z_{nk}\right)$



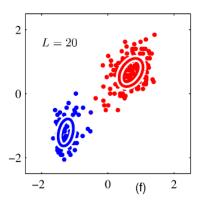
Calculate model parameters $\{\pi_k, oldsymbol{\mu}_k, oldsymbol{\Sigma}_k\}$ using these responsibilities



Iteration 2



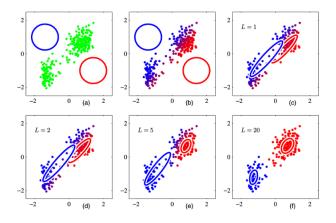
Iteration 5



 $Interation\ 20 - converged$

Gaussian mixture models:d>1

See Bishop (2006) for details



Bayesian Information Criterion (BIC)

- ► How to pick k?
- ► Probabilistic model:

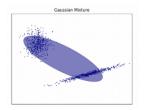
$$\mathbf{L} = \ln p(\mathbf{X} \mid \pi, \mu, \mathbf{\Sigma}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N} \left(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k \right) \right\}$$

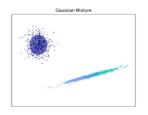
- ► Tries to "fit" the data (maximize likelihood)
- ► Choose *K* that makes L as large as possible?
 - $lackbox{K} = n$: each data point has its own "source"
 - may not work well for new data points
- Split points into training set T and validation set V
 - ► for each k : fit parameters of **T**
 - measure likelihood of V
 - ightharpoonup sometimes still best when k=n
- ▶ "Occam's razor":
 - ▶ Pick the "simplest" of all models that fits the data.
 - Assess, e.g., via Bayes Information Criterion (BIC): $\max_{p} \{L 1/2p * \log(n)\}$
 - L: Likelihood; p: # Parameters in the model how simple is the model.

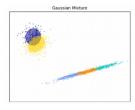
Hands-on example

https://scikit-learn.org/stable/modules/mixture.html demo/GMM_scikit_example.py

- ► Plot the confidence ellipsoids of a mixture of two Gaussians obtained with Expectation Maximization (GaussianMixture class)
- ► The model has access to 1,3, and 5 components with which to fit the data. Note that the Expectation Maximization model will necessarily use ALL components
- ▶ In the 5-component example, we can see that the Expectation Maximization model splits some components arbitrarily, because it is trying to fit too many components.

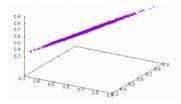


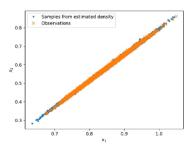




Hands-on example 2

- ➤ We simulate a bunch of data (e.g., an ergodic set). it is in a text file (ergodic_data.txt 3 dimensions)
- ► We apply GMM (build_density.py)
- We can sample data from the fitted GMM model (sample.py)





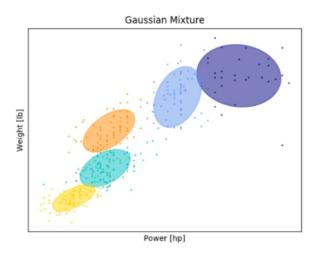
GMM — Cars based on Power and Weight

```
import itertools
import numpy as np
from scipy import linalg
import matplotlib.pyplot as plt
import matplotlib as mpl
from sklearn import mixture
import pandas as pd
from sklearn.model selection import train_test_split
from sklearn.preprocessing import MinMaxScaler
import matplotlib.pyplot as plt
color_iter = itertools.cycle(['navy', 'c', 'cornflowerblue', 'gold', 'darkorange'])
# load data
cars = pd.read_csv('auto-mpg.data.txt',header=None, sep='\s+')
# extract power and weight as data matrix X
X = cars.iloc[:, [3,4]].values
# extract origin as target value y
v = cars.iloc[:, 7].values
# normalize data
min max scaler = MinMaxScaler()
min_max_scaler.fit(X) # determine min and max
X_normalized = min_max_scaler.transform(X)
## Fit a Gaussian mixture with EM using five components
gmm = mixture.GaussianMixture(n_components=5, covariance_type='full').fit(X_normalized)
plot_results(X normalized, gmm.predict(X normalized), gmm.means_, gmm.covariances_, 0,
             'Gaussian Mixture')
plt.show()
```

Cont.

```
def plot_results(X, Y_, means, covariances, index, title):
splot = plt.subplot(1, 1, 1 + index)
for i, (mean, covar, color) in enumerate(zip(
       means, covariances, color_iter)):
    v, w = linalg.eigh(covar)
    v = 2. * np.sqrt(2.) * np.sqrt(v)
    u = w[0] / linalg.norm(w[0])
    # as the DP will not use every component it has access to
    # unless it needs it, we shouldn't plot the redundant
    # components.
    if not np.any(Y_ == i):
        continue
    plt.scatter(X[Y_{=} == i, 0], X[Y_{=} == i, 1], .8, color=color)
    # Plot an ellipse to show the Gaussian component
    angle = np.arctan(u[1] / u[0])
    angle = 180. * angle / np.pi # convert to degrees
    ell = mpl.patches.Ellipse(mean, v[0], v[1], 180. + angle, color=color)
    ell.set_clip_box(splot.bbox)
    ell.set_alpha(0.5)
    splot.add_artist(ell)
plt.xticks(())
plt.yticks(())
plt.xlabel('Power [hp]')
plt.ylabel('Weight [lb]')
plt.title(title)
```

GMM — Cars based on Power and Weight



4. Expectation Maximization: A General Version of EM

▶ In general, we are interested in maximizing the likelihood

$$p(\boldsymbol{X} \mid \boldsymbol{\theta}) = \sum_{\boldsymbol{Z}} p(\boldsymbol{X}, \boldsymbol{Z} \mid \boldsymbol{\theta})$$

where \boldsymbol{X} denotes all observed variables, and \boldsymbol{Z} denotes all latent (hidden, unobserved) variables

- \blacktriangleright Assume that maximizing $p(X \mid \theta)$ is difficult (e.g. mixture of Gaussians)
- ▶ But maximizing $p(X, Z \mid \theta)$ is tractable (everything observed)
 - $ightharpoonup p(X, Z \mid \theta)$ is referred to as the complete-data likelihood function, which we don't have

A Lower Bound

- ► The strategy for optimization will be to introduce a lower bound on the likelihood
 - ► This lower bound will be based on the complete-data likelihood, which is easy to optimize
- Iteratively increase this lower bound
- ► Make sure we're increasing the likelihood while doing so

A Decomposition Trick

▶ To obtain the lower bound, we use a decomposition:

$$\begin{split} \ln p(\boldsymbol{X}, \boldsymbol{Z} \mid \boldsymbol{\theta}) &= \ln p(\boldsymbol{X} \mid \boldsymbol{\theta}) + \ln p(\boldsymbol{Z} \mid \boldsymbol{X}, \boldsymbol{\theta}) \text{ product rule} \\ \ln p(\boldsymbol{X} \mid \boldsymbol{\theta}) &= \mathcal{L}(q, \boldsymbol{\theta}) + \mathcal{K}L(q \parallel p) \\ \mathcal{L}(q, \boldsymbol{\theta}) &\equiv \sum_{\boldsymbol{Z}} q(\boldsymbol{Z}) \ln \left\{ \frac{p(\boldsymbol{X}, \boldsymbol{Z} \mid \boldsymbol{\theta})}{q(\boldsymbol{Z})} \right\} \\ \mathcal{K}L(q \parallel p) &\equiv -\sum_{\boldsymbol{Z}} q(\boldsymbol{Z}) \ln \left\{ \frac{p(\boldsymbol{Z} \mid \boldsymbol{X}, \boldsymbol{\theta})}{q(\boldsymbol{Z})} \right\} \end{split}$$

▶ $\mathrm{KL}(q\|1)$ is known as the Kullback-Leibler divergence (KL-divergence), and is ≥ 0 (next slide) \longrightarrow Hence $\ln p(\boldsymbol{X} \mid \boldsymbol{\theta}) \geq \mathcal{L}(q, \boldsymbol{\theta})$

TODO colors

Kullback-Leibler Divergence

ightharpoonup KL(p(x)||q(x)) is a measure of the difference between distributions p(x) and q(x):

$$KL(p(x)||q(x)) = -\sum_{x} p(x) \log \frac{q(x)}{p(x)}$$

- Motivation: average additional amount of information required to encode x using code assuming distribution q(x) when x actually comes from p(x)
- ▶ Note it is not symmetric: $KL(q(x)||p(x)) \neq KL(p(x)||q(x))$ in general
- ► It is non-negative:
 - ▶ Jensen's inequality: $-\ln\left(\sum_{x} x p(x)\right) \le -\sum_{x} p(x) \ln x$
 - ► Apply to KL:

$$KL(p||q) = -\sum_{x} p(x) \log \frac{q(x)}{p(x)} \ge -\ln \left(\sum_{x} \frac{q(x)}{p(x)} p(x)\right) = -\ln \sum_{x} q(x) = 0$$

Increasing the Lower Bound — E-step

► EM is an iterative optimization technique which tries to maximize this lower bound: $\ln p(X \mid \theta) \ge \mathcal{L}(q, \theta)$

E step: Fix $\pmb{\theta}^{old}$, maximize $\mathcal{L}(q, \pmb{\theta}^{old})$ wrt q i.e. Choose distribution q to maximize \mathcal{L} Reordering bound:

$$\mathcal{L}(q, \boldsymbol{\theta}^{old}) = \ln p(\boldsymbol{X}|\boldsymbol{\theta}^{old}) - KL(q||p)$$

 $\ln p(\boldsymbol{X}|\boldsymbol{\theta}^{old})$ does not depend on qMaximum is obtained when KL(q||p) is as small as possible Occurs when q=p, i.e. $q(\boldsymbol{Z})=p(\boldsymbol{Z}|\boldsymbol{X},\boldsymbol{\theta})$ This is the posterior over \boldsymbol{Z} , recall these are the responsibilities from MoG model

Increasing the Lower Bound — M-step

M step: Fix q, maximize $\mathcal{L}(q, \theta)$ wrt θ The maximization problem is on

$$\begin{array}{c} \mathcal{L}(q,\theta) = \sum_{\textbf{\textit{Z}}} q(\textbf{\textit{Z}}) \ln p(\textbf{\textit{X}},\textbf{\textit{Z}} \mid \theta) - \sum_{\textbf{\textit{Z}}} q(\textbf{\textit{Z}}) \ln q(\textbf{\textit{Z}}) \\ = \sum_{\textbf{\textit{Z}}} p\left(\textbf{\textit{Z}} \mid \textbf{\textit{X}}, \theta^{\text{old}}\right) \ln p(\textbf{\textit{X}},\textbf{\textit{Z}} \mid \theta) - \sum_{\textbf{\textit{Z}}} p\left(\textbf{\textit{Z}} \mid \textbf{\textit{X}}, \theta^{\text{old}}\right) \ln p\left(\textbf{\textit{Z}} \mid \textbf{\textit{X}}, \theta^{\text{old}}\right) \end{array}$$

Second term is constant with respect to θ First term is In of complete data likelihood, which is assumed easy to optimize Expected complete log likelihood - what we think complete data likelihood will be

Why does EM work?

- ▶ In the M-step we changed from θ^{old} to θ^{new}
- ► This increased the lower bound *L*, unless we were at a maximum (so we would have stopped)
- ► This must have caused the log likelihood to increase
- ► The E-step set q to make the KL-divergence 0:

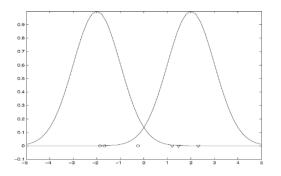
$$\ln p\left(\boldsymbol{X} \mid \boldsymbol{\theta}^{\mathsf{old}} \right) = \mathcal{L}\left(q, \boldsymbol{\theta}^{\mathsf{old}} \right) + \mathit{KL}(q \| p) = \mathcal{L}\left(q, \boldsymbol{\theta}^{\mathsf{old}} \right)$$

lacktriangle Since the lower bound L increased when we moved from $heta^{
m old}$ to $heta^{
m new}$

$$\ln p\left(\boldsymbol{X} \mid \boldsymbol{\theta}^{\mathsf{old}}\right) = \mathcal{L}\left(q, \boldsymbol{\theta}^{\mathsf{old}}\right) < \mathcal{L}\left(q, \boldsymbol{\theta}^{\mathsf{new}}\right) \\
= \ln p\left(\boldsymbol{X} \mid \boldsymbol{\theta}^{\mathsf{new}}\right) - KL\left(q \| p^{\mathsf{new}}\right)$$

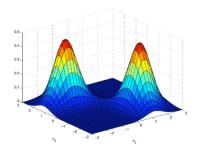
lacktriangle So the log-likelihood has increased going from $heta^{
m old}$ to $heta^{
m new}$

Bounding Example

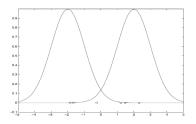


Consider 2 component 1-D MoG with known variances.

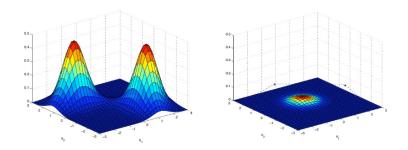
Bounding Example



True likelihood function Recall we're fitting means θ_1, θ_2



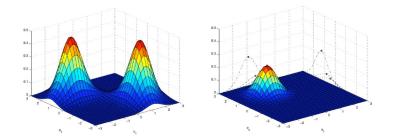
Bounding Peaks



- \blacktriangleright Lower bound the likelihood function using averaging distribution q(Z)

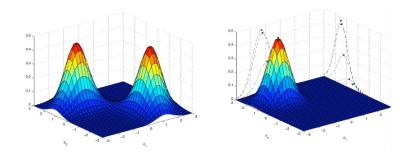
 - ▶ In $p(\mathbf{X} \mid \theta) = \mathcal{L}(q, \theta) + \mathcal{K}L(q(\mathbf{Z}) || p(\mathbf{Z} \mid \mathbf{X}, \theta))$ ▶ Since $q(\mathbf{Z}) = p(\mathbf{Z} \mid \mathbf{X}, \theta^{\text{old}})$, bound is tight (equal to actual likelihood) at $\theta = \theta^{\text{old}}$

Bounding Peaks



- ightharpoonup Lower bound the likelihood function using averaging distribution q(Z)
 - $\blacktriangleright \ln p(\boldsymbol{X} \mid \boldsymbol{\theta}) = \mathcal{L}(q, \boldsymbol{\theta}) + KL(q(\boldsymbol{Z}) || p(\boldsymbol{Z} \mid \boldsymbol{X}, \boldsymbol{\theta}))$
 - Since $q(\mathbf{Z}) = p(\mathbf{Z} \mid \mathbf{X}, \theta^{\text{old}})$, bound is tight (equal to actual likelihood) at $\theta = \theta^{\text{old}}$

Bounding Peaks



- ightharpoonup Lower bound the likelihood function using averaging distribution q(Z)
 - $\blacktriangleright \ln p(\boldsymbol{X} \mid \boldsymbol{\theta}) = \mathcal{L}(q, \boldsymbol{\theta}) + KL(q(\boldsymbol{Z}) || p(\boldsymbol{Z} \mid \boldsymbol{X}, \boldsymbol{\theta}))$
 - Since $q(\mathbf{Z}) = p(\mathbf{Z} \mid \mathbf{X}, \theta^{\text{old}})$, bound is tight (equal to actual likelihood) at $\theta = \theta^{\text{old}}$

Recall About the EM Algorithm

Some good things about EM:

- ▶ no learning rate (step-size) parameter.
- automatically enforces parameter constraints.
- very fast for low dimensions.
- each iteration guaranteed to improve likelihood.

Some bad things about EM:

- ► can get stuck in local minima.
- can be slower than conjugate gradient (especially near convergence).
- requires expensive inference step.
- is a maximum likelihood/MAP (maximum a posterior) method.

EM — Summary

► EM finds local maximum to likelihood

$$p(\boldsymbol{X} \mid \boldsymbol{\theta}) = \sum_{\boldsymbol{Z}} p(\boldsymbol{X}, \boldsymbol{Z} \mid \boldsymbol{\theta})$$

- ► Iterates two steps:
 - ► E step "fills" in the missing variables Z (calculates their distribution)
 - M step maximizes expected complete log likelihood (expectation wrt E step distribution)
- ► This works because these two steps are performing a coordinatewise hill-climbing on a lower bound on the likelihood $p(X | \theta)$

5. Hierarchical Clustering

- ► k-Means determines a flat clustering of data points; there is no relationship between the clusters
- ► Hierarchical clustering determines a sequence of increasingly fine-grained clusterings

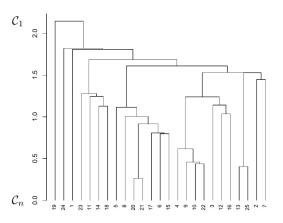
$$C_1,\ldots,C_n$$

- $lackbox{}{\mathcal{C}}_1 = \{\mathcal{D}\}$ contains all data points in a single cluster
- ▶ $C_n = \{\{x_i\} : x_i \in \mathcal{D}\}$ contains one cluster per data point
- ► Clustering C_i is contained in clustering C_{i-1}

$$\forall C_j \in C_i : \exists C_l \in C_{i-1} : C_j \subseteq C_l$$

Dendrogram

Sequence of clusterings can be visualized in a dendrogram



Hierarchical Agglomerative vs. Divisive Clustering

- ► Hierarchical Agglomerative Clustering (HAC)
 - ightharpoonup starts with the most fine-grained clustering \mathcal{C}_n
 - ▶ proceeds bottom-up and merges the two closest clusters in C_i to obtain the more coarse-grained clustering C_{i-1}
- ► Hierarchical Divisive Clustering (HDC)
 - lacktriangle starts with the most coarse-grained clustering \mathcal{C}_1
 - ightharpoonup proceeds top-down and splits one of the clusters in \mathcal{C}_{i-1} to obtain the more fine-grained clustering \mathcal{C}_1

Hierarchical Agglomerative vs. Divisive Clustering

- ► Hierarchical Agglomerative Clustering (HAC)
 - ightharpoonup starts with the most fine-grained clustering C_n
 - ightharpoonup proceeds bottom-up and merges the two closest clusters in C_i to obtain the more coarse-grained clustering C_{i-1}
- ➤ So far, we can only measure distance between data points, but we need a measure of distance between clusters

Linkage Criteria

Linkage criteria measure distance between two clusters based on the distance between data points therein Single-Link

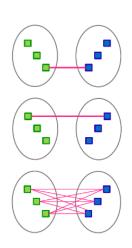
$$\delta(C_i, C_j) = \min \{d(\mathbf{x}, \mathbf{y}) \mid \mathbf{x} \in C_i, \mathbf{y} \in C_j\}$$

Complete-Link

$$\delta(C_i, C_j) = \max\{d(\mathbf{x}, \mathbf{y}) \mid \mathbf{x} \in C_i, \mathbf{y} \in C_j\}$$

Average-Link

$$\delta(C_i, C_j) = \frac{1}{|C_i| |C_j|} \sum_{\mathbf{x} \in C_i} \sum_{\mathbf{y} \in C_i} d(\mathbf{x}, \mathbf{y})$$



Pseudocode: Hierarchical Agglomerative Clustering

```
// Start with each data point in a separate cluster  \mathcal{C}_n = \big\{ \{\mathbf{x}_i\} \, : \, \mathbf{x}_i \in \mathcal{D} \big\};  for (int t = n; t > 1; t--) {    // Determine the two clusters closest to each other  C_i^*, C_j^* = \underset{C_i, C_j \in \mathcal{C}_t \, : \, C_i \neq C_j}{\arg\min} \; \delta(C_i, C_j);  // Merge the two clusters  \mathcal{C}_{t-1} = \big( \mathcal{C}_t \setminus \big\{ C_i^*, C_j^* \big\} \big) \cup \big\{ C_i^* \cup C_j^* \big\};  }
```

HAC Example

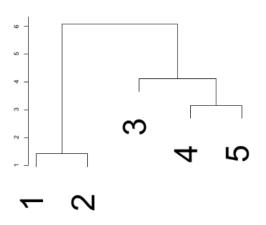
ightharpoonup Consider the following data points in \mathbb{R}^2

$$\begin{array}{l} \textbf{x_1} = (1,0) \\ \textbf{x_2} = (2,1) \\ \textbf{x_3} = (8,0) \\ \textbf{x_4} = (12,1) \\ \textbf{x_5} = (15,1) \end{array} \quad \textbf{d} = \left[\begin{array}{ccccc} 0.00 & 1.41 & 7.00 & 11.05 & 14.04 \\ 0.00 & 6.08 & 10.00 & 13.04 \\ 0.00 & 4.12 & 7.07 \\ 0.00 & 3.00 \\ 0.00 \end{array} \right]$$

► With distance matrix d

HAC with Single-Link Example

HAC with single-link based on distance matrix d



$$\mathcal{C}_1 = \left\{ \left\{ \mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3, \mathbf{x}_4, \mathbf{x}_5 \right\} \right\}$$

$$\mathcal{C}_2 = \left\{ \left\{ \mathbf{x}_1, \mathbf{x}_2 \right\}, \left\{ \mathbf{x}_3, \mathbf{x}_4, \mathbf{x}_5 \right\} \right\}$$

$$\mathcal{C}_3 = \left\{ \left\{ \mathbf{x}_1, \mathbf{x}_2 \right\}, \left\{ \mathbf{x}_3 \right\}, \left\{ \mathbf{x}_4, \mathbf{x}_5 \right\} \right\}$$

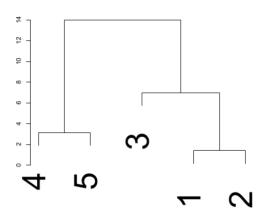
$$\mathcal{C}_4 = \left\{ \left\{ \mathbf{x}_1, \mathbf{x}_2 \right\}, \left\{ \mathbf{x}_3 \right\}, \left\{ \mathbf{x}_4 \right\}, \left\{ \mathbf{x}_5 \right\} \right\}$$

$$\mathcal{C}_5 = \left\{ \left\{ \mathbf{x}_1 \right\}, \left\{ \mathbf{x}_2 \right\}, \left\{ \mathbf{x}_3 \right\}, \left\{ \mathbf{x}_4 \right\}, \left\{ \mathbf{x}_5 \right\} \right\}$$

$$\mathbf{d} = \begin{bmatrix} 0.00 & 1.41 & 7.00 & 11.05 & 14.04 \\ 0.00 & 6.08 & 10.00 & 13.04 \\ 0.00 & 4.12 & 7.07 \\ 0.00 & 3.00 \\ 0.00 & 0.00 \end{bmatrix}$$

HAC with Single-Link Example

HAC with complete-link based on distance matrix d



$$C_2 = \{\{\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3\}, \{\mathbf{x}_4, \mathbf{x}_5\}\}$$

$$C_3 = \{\{\mathbf{x}_1, \mathbf{x}_2\}, \{\mathbf{x}_3\}, \{\mathbf{x}_4, \mathbf{x}_5\}\}$$

$$C_4 = \{\{\mathbf{x}_1, \mathbf{x}_2\}, \{\mathbf{x}_3\}, \{\mathbf{x}_4\}, \{\mathbf{x}_5\}\}$$

$$C_5 = \{\{\mathbf{x}_1\}, \{\mathbf{x}_2\}, \{\mathbf{x}_3\}, \{\mathbf{x}_4\}, \{\mathbf{x}_5\}\}$$

$$0.00 \quad 1.41 \quad 7.00 \quad 11.05 \quad 14.04 \quad 7.00 \quad 13.04 \quad 7.00 \quad 13.04 \quad 7.00 \quad 13.04 \quad 7.00 \quad 9.00 \quad 3.00 \quad 9.00 \quad 3.00$$

 $C_1 = \{\{\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3, \mathbf{x}_4, \mathbf{x}_5\}\}$

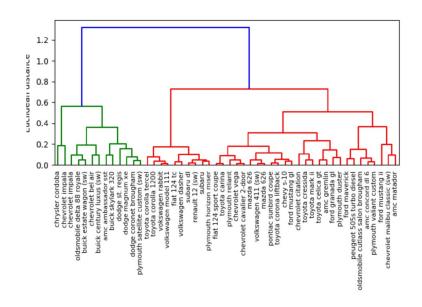
0.00

Clustering Cars based on Power and Weight

demo/HAC_example.py

```
import numpy as np
import pandas as pd
from sklearn.preprocessing import MinMaxScaler
from scipy.cluster.hierarchy import linkage
from scipy.cluster.hierarchy import dendrogram
import matplotlib.pyplot as plt
# load data
cars = pd.read csv('auto-mpg.data.txt',header=None, sep='\s+')
# keep a sample of 50 cars
cars = cars.sample(50, random_state=0)
# extract labels
labels = cars.iloc[:,8].values
# extract power and weight as data matrix X
X = cars.iloc[:, [3,4]].values
# normalize data
min max scaler = MinMaxScaler()
min_max_scaler.fit(X) # determine min and max
X normalized = min max scaler.transform(X)
# perform hierarchical agglomerative clustering using complete linkage
clusters = linkage(X normalized, method='complete', metric='euclidean')
# plot dendrogram
dendrogram = dendrogram(clusters, labels=labels)
plt.tight lavout()
plt.ylabel('Euclidean distance')
plt.show()
```

Clustering Cars based on Power and Weight



6. Density-based Clustering

- ► k-Means as a representative-based clustering method can only find convey clusters and must assign every data point to a cluster.
- ► Density-based clustering methods determine clusters as regions having consistently high density and label isolated data points as noise
- ► Density-Based Spatial Clustering of Applications with Noise (DBSCAN)

Density-Based Clustering

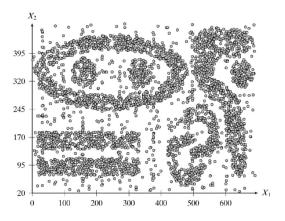


Fig. From Zaki and Meira (2014)

DBSCAN — the idea

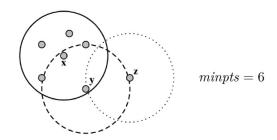
► Epsilon Neighborhood of a data point *x*

$$N_{\epsilon}(\mathbf{x}) = \{ \mathbf{y} \mid d(\mathbf{x}, \mathbf{y}) \leq \epsilon \}$$

contains all points having distance less than or equal to arepsilon

- ▶ Data point x is called a core to ε its epsilon neighborhood contains at least minpts data points (including x)
- ▶ Data point x is called a border point, if it is not a core, but belongs to the epsilon neighborhood of a core
- ► All other data points are considered noise

Core, Border, and Noise



- ▶ Data point x is a core
- ► Data point y is a border point
- ► Data point *z* is noise

Reachability

▶ Data point *x* is directly reachable from data point *y*, if *y* is a core and *x* belongs to the epsilon neighborhood of *y*, i.e.

$$\mathbf{x} \in \mathit{N}_{\epsilon}(\mathbf{y})$$

- ▶ Data point x is (density) reachable from data point y,
- ightharpoonup if there is a chain of data points x_0, \ldots, x_1 , so that

$$x_0 = x \wedge x_l = y$$

 $\forall 1 \leq i \leq l : x_i$ is directly reachable from x_{i-1}

► Reachability is not symmetric, since the data point y could be a core, but the data point x is not

Connectedness and Density-Based Clusters

- ► Two data points x and y are called connected, if there is a core z, so that both x and y are reachable from z
- ▶ Density-based cluster is a maximal subset of connected data points, i.e., there are no data points that could be added

DBSCAN

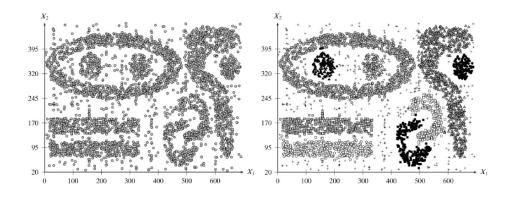
- Intuition:
 - Compute epsilon neighborhoods for all data points
 - ► Determine all cores
 - ► Determine noise
 - ► Grow a new density-based cluster from each data point that does not yet belong to an already-determined cluster
- ► Note that DBSCAN is not deterministic, since the assignment of data point to clusters depends on the order in which data points are considered

Pseudo-code DBSCAN

```
dbScan(D, \epsilon, minpts) {
  // Cores
   Cores = \emptyset;
  for (x \in \mathcal{D}) {
     // Compute epsilon neighborhoods
     N_{\epsilon}(\mathbf{x}) = \text{computeNeighborhood}(\mathbf{x}, \epsilon);
     // Initialize cluster id
     id(\mathbf{x}) = \emptyset;
     // Check whether data point is a core
     if (N_{\epsilon}(\mathbf{x}) \geq minpts) Cores = Cores \cup \{\mathbf{x}\};
   // Grow density-based cluster from each core
  k = 0:
  for (x ∈ Cores) {
     if(id(x) == \emptyset) {
         k++;
         id(x) = k:
         densityConnected(x,k);
  // Determine clustering, border points, and noise
  C = \emptyset:
  for (i = 1...k) C = C \cup \{\{\mathbf{x} \in \mathcal{D} : id(x) = k\}\};
   Noise = \{ \mathbf{x} \in \mathcal{D} : id(\mathbf{x}) = \emptyset \} ;
   Border = D \setminus \{Cores \cup Noise\}:
```

```
\begin{array}{l} \texttt{densityConnected}(\mathbf{x},\ k)\ \{\\ \texttt{for}\left(\mathbf{y}\in N_{\epsilon}(\mathbf{x})\right)\ \{\\ id(\mathbf{y})=k\,;\\ \texttt{if}\ (\mathbf{y}\in \mathit{Cores})\ \texttt{densityConnected}(\mathbf{y},\ k)\,;\\ \}\\ \} \end{array}
```

DBSCAN in Action

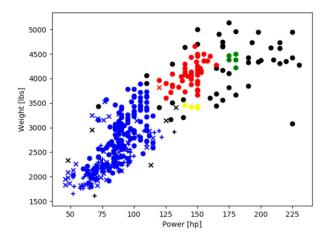


Clustering Cars based on Power and Weight

demo/DBSCAN_example.py

```
import numpy as np
import pandas as pd
from sklearn.model selection import train_test_split
from sklearn.preprocessing import MinMaxScaler
from sklearn.cluster import DBSCAN
from matplotlib.backends.backend pdf import PdfPages
import matplotlib.pyplot as plt
# load data
cars = pd.read_csv('auto-mpg.data.txt',header=None, sep='\s+')
# extract power and weight as data matrix X
X = cars.iloc[:, [3.4]].values
# extract origin as target value y
y = cars.iloc[:, 7].values
# normalize data
min_max_scaler = MinMaxScaler()
min_max_scaler.fit(X) # determine min and max
X normalized = min max scaler.transform(X)
# DRSCAN
db = DBSCAN(eps=0.05,min_samples=5,metric='euclidean')
db.fit_predict(X_normalized)
# plot cars
# U.S. : o / Europe: x / Japan : +
m = ['o' \text{ if } o == 1 \text{ else 'x' if } o == 2 \text{ else '+' for } o \text{ in } v]
# Noise : black / Cluster 1 : red / Cluster 2 : blue /
# Cluster 3 : green / Cluster 4 : yellow
c = ['black' if l==-1 else 'red' if l==0 else 'blue' if l==1
else 'green' if 1==2 else 'vellow' for 1 in db.labels ]
for i in range(0,len(X)):
    plt.scatter(X[i,0], X[i,1], color=c[i], marker=m[i])
plt.xlabel('Power [hp]')
plt.ylabel('Weight [lbs]')
plt.show()
```

Clustering Cars based on Power and Weight



Summary

- ► Hierarchical clustering determines a sequence of clusterings that can be visualized in a dendrogram
- ► DBSCAN as a density-based clustering method can find non-convex clusters and is able to label data points as noise
- ▶ DBSCAN comes with two hyper parameters ε and minpts that need to be carefully tuned based on the data

